

# The Sample Complexity of Semi-Supervised Learning with Nonparametric Mixture Models

Chen Dan, Liu Leqi, Bryon Aragam, Pradeep Ravikumar, Eric Xing  
Carnegie Mellon University

## Overview

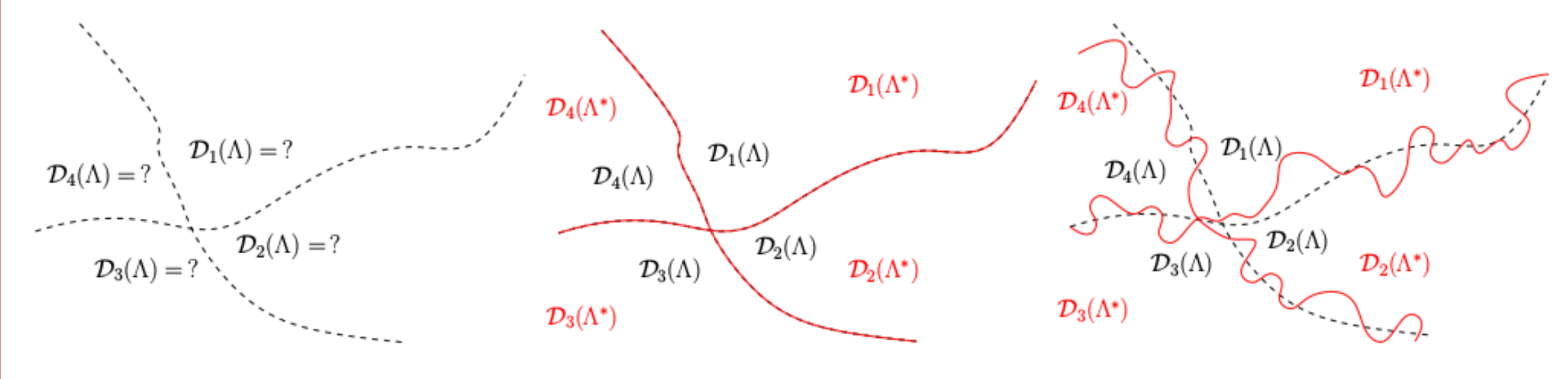
- A novel framework for analyzing the sample complexity of semi-supervised learning (SSL) in general, nonparametric settings.
- Establish  $\Omega(K \log K)$  sample complexity for learning the class assignment and provide conditions under which the resulting classifier converges to the Bayes classifier.
- Provide efficient algorithms in learning the class assignment and illustrate their performance on real and simulated data.

## SSL as Permutation Learning

Let  $K$  be the number of classes in the output space  $\mathcal{Y}$ . We formulate SSL as follows (see Figure below):

1. Use the unlabeled data to learn a  $K$ -component nonparametric mixture model  $\Lambda$  that approximates the unlabeled data density  $F^*$ ;
2. Use the labeled data to learn an assignment  $\pi : [K] \rightarrow \mathcal{Y}$  between decision regions  $\mathcal{D}_b(\Lambda)$  and classes  $\alpha_k$ ;
3. Given a test point  $X$ , first assign to a decision region, then use  $\pi$  to assign a label.

The pair  $(\Lambda, \pi)$  thus defines a classifier  $g_{\Lambda, \pi} : \mathcal{X} \rightarrow \mathcal{Y}$  that we analyze.



## Assumptions

We make the following general assumptions:

1. **Nonparametric.** The class conditional distributions  $\mathbb{P}(X | Y)$  can be anything, and we do not assume any parametric form.
2. **Multi-class ( $K > 2$ ).** Existing techniques based on Neyman-Pearson classification no longer apply.
3. **Unknown  $\mathbb{P}(X), \mathbb{P}(X | Y)$ .** Both the unlabeled data distribution and class conditionals are unknown and unidentified.

These assumptions generalize existing work, which either assume  $K = 2$  or that the unlabeled data distribution is either known, or approximately known.

## References

- Castelli, V. and Cover, T. M. (1996), IEEE Trans. Inform. Theory 42(6): 2102-2117.
- Aragam, B., Dan, C., Ravikumar, P., and Xing, E.P. (2018), arXiv: 1802.04397.
- Rigollet, P. (2007), JMLR 8(Jul): 1369-1392.
- Singh, A., Nowak, R. and Zhu, X. (2009), NeurIPS: 1513-1520.

## Sample Complexity

**Maximum likelihood.** The maximum likelihood estimator (MLE) is given by:

$$\hat{\pi}_{\text{MLE}} \in \operatorname{argmax}_{\pi} \ell_n(\pi; \Lambda), \quad \ell_n(\pi; \Lambda) := \frac{1}{n} \sum_{i=1}^n \log \lambda_{\pi(Y^i)} f_{\pi(Y^i)}(X^i).$$

Given  $\Lambda$ , the notation  $\mathbb{E}_* \ell(\pi; \Lambda, X, Y) = \mathbb{E}_* \log \lambda_{\pi(Y)} f_{\pi(Y)}(X)$  denotes the expectation of the *misspecified* log-likelihood with respect to the *true* distribution. Define the “gap”

$$\Delta_{\text{MLE}}(\Lambda) := \mathbb{E}_* \ell(\pi^*; \Lambda, X, Y) - \max_{\pi \neq \pi^*} \mathbb{E}_* \ell(\pi; \Lambda, X, Y). \quad (1)$$

For any function  $a : \mathbb{R} \rightarrow \mathbb{R}$ , define the usual Fenchel-Legendre dual  $a^*(t) = \sup_{s \in \mathbb{R}} (st - a(s))$ . Let  $U_b = \log \lambda_b f_b(X)$  and  $\beta_b(s) = \log \mathbb{E}_* \exp(sU_b)$ .

**Theorem 1** (Sample complexity of MLE). *Suppose that  $\lambda_k^* = 1/K$  for each  $k$ ,  $\Delta_{\text{MLE}} > 0$ , and*

$$n \geq K \log(K/\delta) \left[ 1 + \frac{4}{\inf_b \beta_b^*(\Delta_{\text{MLE}}/3)} \right].$$

Then  $\mathbb{P}(\hat{\pi}_{\text{MLE}} = \pi^*) \geq 1 - \delta$ .

**Majority vote.** The majority vote estimator (MV) is given by a simple majority vote over amongst the labels in each decision region. For any  $\Lambda$ , define  $m_b := |i : X^{(i)} \in \mathcal{D}_b(\Lambda)|$  and  $\chi_{bj}(\Lambda) := \frac{1}{m_b} \sum_{i=1}^n \mathbf{1}(Y^{(i)} = j, X^{(i)} \in \mathcal{D}_b(\Lambda))$ , where  $\mathbf{1}(\cdot)$  is the indicator function. Similar to the MLE, our results for MV depend crucially on a “gap” quantity, given by

$$\Delta_{\text{MV}}(\Lambda) := \inf_b \left\{ \mathbb{E}_* \chi_{bb}(\Lambda) - \max_{j \neq b} \mathbb{E}_* \chi_{bj}(\Lambda) \right\}. \quad (2)$$

**Theorem 2** (Sample complexity of MV). *Suppose that  $\mathbb{P}(X \in \mathcal{D}_b(\Lambda)) = 1/K$  for each  $k$ ,  $\Delta_{\text{MV}} > 0$ , and*

$$n \geq K \log(K/\delta) \left[ 1 + \frac{18}{\Delta_{\text{MV}}^2} \right].$$

Then  $\mathbb{P}(\hat{\pi}_{\text{MV}} = \pi^*) \geq 1 - \delta$ .

## Classification Error

We can further bound the classification error of the classifier in terms of the Wasserstein distance  $W_1(\Lambda, \Lambda^*)$  between  $\Lambda$  and  $\Lambda^*$  as follows:

**Theorem 3** (Classification error). *Let  $g^* = g_{\Lambda^*, \pi^*}$  denote the Bayes classifier. If  $\pi^*(\alpha_b) = \operatorname{argmin}_i d_{\text{TV}}(f_i, f_b^*)$  then there is a constant  $C > 0$  depending on  $K$  and  $\Lambda^*$  such that*

$$\mathbb{P}(g_{\Lambda, \pi^*}(X) \neq Y) \leq \mathbb{P}(g^*(X) \neq Y) + C \cdot W_1(\Lambda, \Lambda^*) + \sum_b |\lambda_{\pi^*(\alpha_b)} - \lambda_b^*|.$$

This theorem allows for the possibility that the mixture model  $\Lambda$  learned from the unlabeled data is not the same as  $\Lambda^*$ . It is thus necessary to assume that the mismatch between  $\Lambda$  and  $\Lambda^*$  is not so bad that the closest density  $f_i$  to  $f_b^*$  is something other than  $f_{\pi^*(\alpha_b)}$ .

## Algorithms

Define  $C_k = \{i : Y^{(i)} = \alpha_k\}$ .

**MLE.** The MLE can be found via the Hungarian algorithm, by exploiting a connection with max weight perfect matching in the bipartite graph  $G = (V_{K,K}, w)$  with  $w(k, k') = \sum_{i \in C_k} \log(\lambda_{k'} f_{k'}(X^{(i)}))$ .

**Majority vote.** This is straightforward to compute.

**Greedy.** Assign the  $k$ th class to  $\hat{\pi}_G(\alpha_k) = \operatorname{argmax}_{k' \in [K]} w(k, k') = \operatorname{argmax}_{k' \in [K]} \sum_{i \in C_k} \log(\lambda_{k'} f_{k'}(X^{(i)}))$ . This greedy heuristic can be viewed as a “soft interpolation” of  $\hat{\pi}_{\text{MLE}}$  and  $\hat{\pi}_{\text{MV}}$ .

## Performance of the Algorithms

To test the performance of the three algorithms, we consider three settings: (i) Mixtures of Gaussians, (ii) A nonparametric mixture model, and (iii) MNIST.  $\mathbb{P}(\hat{\pi} = \pi^*)$  is evaluated in two settings: (i)  $\Lambda = \Lambda^*$  and (ii)  $\Lambda \neq \Lambda^*$ .

In terms of classification accuracy, each algorithm was compared with a canonical supervised baseline for MNIST, the LeNet convolutional neural network. As shown below, all three estimators attain higher accuracy with fewer labeled samples, but the accuracy plateaus around 95% since  $\Lambda \neq \Lambda^*$ .

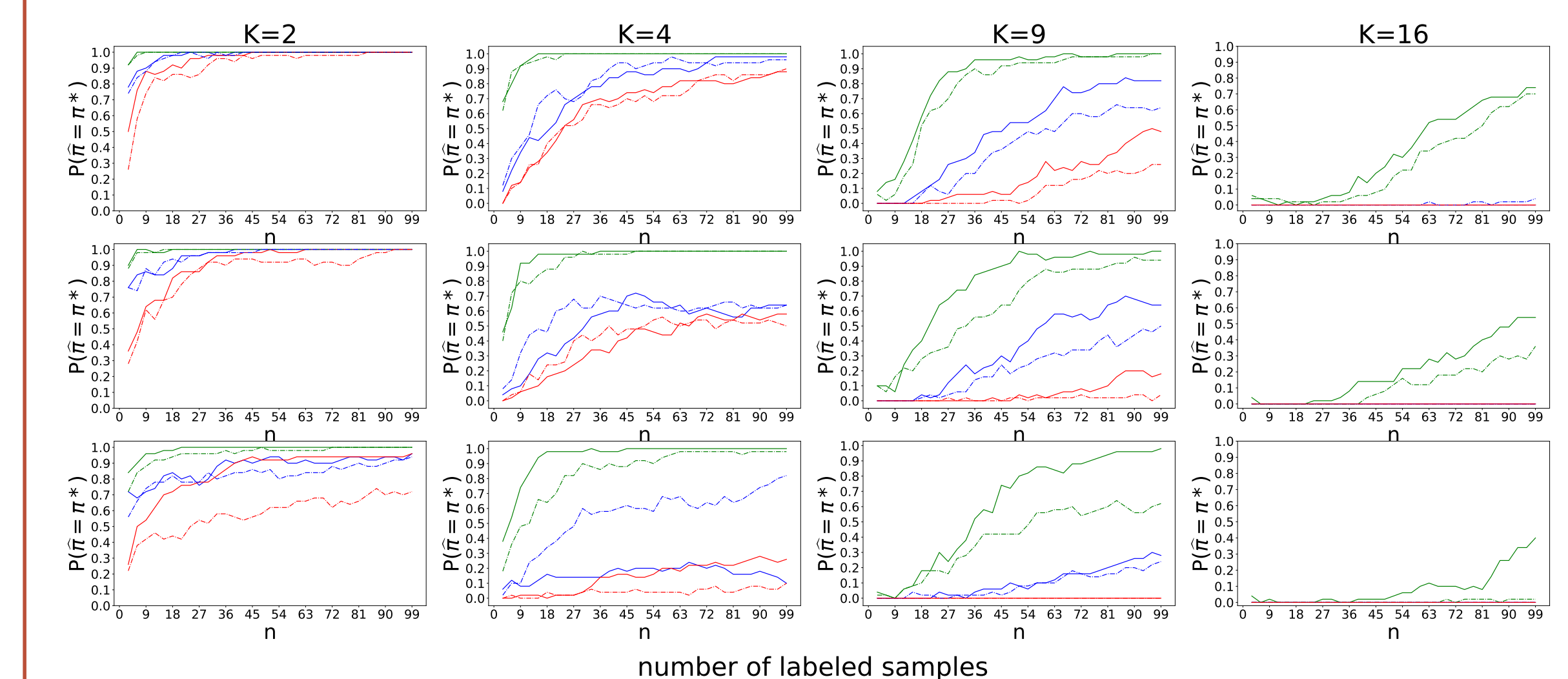


Figure 1

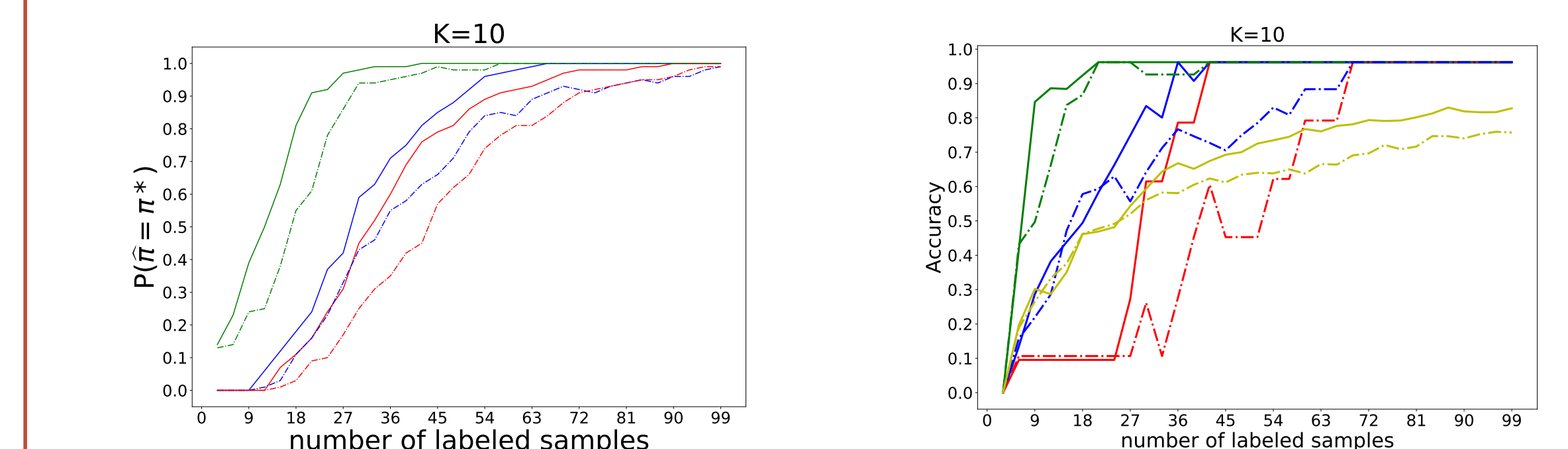


Figure 2

Figure 3

Figure 1 (Mixture of Gaussian) and Figure 2 (MNIST) report performance of MLE (Hungarian - Green; Greedy - Blue) and MV (Red) on learning the class assignment. Solid line and dashed line correspond to the performance when  $\Lambda^* = \Lambda$  and  $\Lambda^* \neq \Lambda$ , respectively. In Figure 1, columns correspond to the number of classes  $K$ ; rows correspond to decreasing separation; e.g. the bottom rows in each figure are the least separated. Figure 3 shows classification accuracy of the SSL estimators and LeNet (Yellow) on MNIST.